Internet Appendix for "Exploiting a Rare Communication Shift to Document the Persuasive Power of News Media"

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Figure 1S: Front Page of the *Sun* on Election Day, May 1, 1997



Table 1S: Coding of Variables

Variable	Coding	BEPS Variable
	(unless otherwise specified, all variables coded to	Label
	vary between 0 and 1 and measured in 1992)	
Vote Choice in 1997	Indicator variable for Labour Party vote in 1997	
(dependent variable)		vote97
	Indicator variable for whether the respondent reads a	
	daily morning paper and the paper he or she reports	
Treatment	reading most often is the Sun, Daily Star,	
	Independent, or Financial Times	readpa96, whpapr96
Prior Labour Vote	Indicator variable for Labour Party vote in 1992.	vote92
Prior Conservative	Indicator variable for Conservative Party vote in 1992.	
Vote		
Prior Liberal Vote	Indicator variable for Liberal Party vote in 1992.	
Prior Labour Party	Indicator variable for identification with the Labour	prtyid92
Identification	Party.	
Prior Conservative	Indicator variable for identification with the	prtyid92
Party Identification	Conservative Party.	
Prior Liberal Party	Indicator variable for identification with the Liberal	prtyid92
Identification	Party.	
	Question asking whether they favor or oppose the	
Prior Labour Party	Labour Party on a 7-category scale, ranging from	labfel92
Support	"strongly favor" to strongly oppose."	
Prior Conservative	Question asking whether they favor or oppose the	confel92
Party Support	Conservative Party on a 7-category scale, ranging from	
	"strongly favor" to strongly oppose."	
Prior Political	10-item index of factual knowledge constructed by the	polqiz92
Knowledge	BEPS investigators categorized into terciles.*	
Prior Television	An index of television news viewership based on how	wtbbc992,
Viewer	often respondents watched news programs on BBC1,	wtbbc692,
	BBC2, ITN, and Chanel 4. For each program,	wtbbcn92, wtitnx92
	respondents could answer never watching, once a week	wtitn592, wtitn592,
	or less, two to three days a week, and four or more days	wtc4, newsat92
	a week.	
Prior Read Daily	Indicator variable for whether the respondent reads any	whpapr92
Newspaper Reader	newspaper. We do not match on this variable (so as not	
	to exclude nonreaders from the control group) but	
	include it in the probit models.	
Prior Ideology	An index constructed with a series of 6 policy	airsh92, richlw92
	preference questions. Higher values indicate a more	tuntnd92, priven92
	leftist ideology and lower values a more rightist	publco92, govrsp92
Duriou Idealesteri	ideology.	Como os s1
Prior Ideological	Created by "folding over" the 6-item ideology scale so	Same as above.
Moderation	that scores of 1 correspond with scores of .5 on the	
	original scale and scores of 0 correspond with Ideology score of 0 or 1.	
Prior		tradula concora
Authoritarianism	An index of 6 policy preference questions. Higher values indicate more authoritarianism.	tradv192, censor92 pubmee92,
Aunoriarianism	values multate more autiontariamism.	publice92,

		gaysex92 tolern92,
		banpty92
Prior Trade Union	Indicator variable for whether the respondent has	tusa92
Member	someone in his/her family who has, at some point, been	
	a member of a trade union.	
Prior Working-	Indicator variable for whether respondent self-	srsoc192 or
Class Identification	identified as working class.	srsoc292
Parents Voted	Indicator variable for whether the respondent	mumvot92 and
Labour	remembers one or both parents voting for Labour when	dadvot92
	he or she was a child.	
Prior Coping with	Treated as a continuous variable with four categories:	copemg92
Mortgage	Mortgage Very Difficult (1), Mortgage a Bit Difficult	•••p•····8>=
1120118480	(0.5), and Not Really Difficult or No Mortgage (0).	
	Separated into five categories: <i>Less than O level</i> (or	hedqul92
Prior Education	foreign qualifications) (0), O Level or Equivalent (.25),	neuquiyz
Thor Education	A Level or Equivalent (.5), Some Higher Education	
	(.75), and College Degree (1).*	
Prior Income	(.75), and Conege Degree (1). £5999 or Less (0), £6000-£11,999 (.33), £12,000-	hhincq92
1 nor meome	$\pm 19,999$ (.67), and $\pm 20,000$ or More (1).*	mmeq 92
Duion Ago	18-24 (0), 25-34 (.167), 35-44 (.333), 45-54 (.5), 55-59	ragect92
Prior Age	(.667), 60-64 (.833), 65+(1), and Prior Age Not	Tagect92
	Given.*	
Cardan		
Gender	Indicator variable for male.	rsex92
White	Indicator variable for white ethnic identity	raceor92
Prior Profession	Indicator variable for the following profession	rsegg92
	categories are included in the probit models:	
	Unemployed, Employer/Manager, Professional, Non-	
D ' D '	Manual, Personal Service, Manual, and Other.†	
Prior Region	Indicator variable (fixed effects) for the following	streg92
	region categories are included in the probit models:	
	North, Northwest, Yorks and Humberside, West	
	Midlands, East Midlands, East Anglia, Southwest	
	England, Southeast England, Greater London, Wales,	
	and Scotland.†	
Prior European	10-point scale running from "feel[ing] that Britain	recind96
Integration Views	should do all it can to unite fully with the European	
	Union (sometimes still called the European	
	Community)" (1) to "feel[ing] that Britain should do all	
	it can to protect its independence from the European	
	Union" (0).‡	
1996 Economy	Indicator variable based on a retrospective question	ec1ok96w
	from the 1996 wave that ranges from "got a lot weaker"	
	(0) to "got a lot stronger" (1).‡	
		1 1 11

* In the matching procedures, treated as a continuous variable. In the probit models, indicator variables are used for the response categories.[†] Not incorporated in matching, but indicator variables are used for response categories in the probit models. [‡] Used only in probit models with 1996 as the baseline.

Table 2S: Measuring the Treatment in 1992 (instead of 1996)

A potential concern we address in the paper is that readers may have sensed the *Sun*'s (or other switching papers) shift to labor in advance. To address this, we present an instrumental variables analysis in the paper (and present more below) that instruments the treatment, measured in 1996, with readership in 1992. Since instrumental variables estimates are exceedingly sensitive to violations of the exclusion restriction, we also present here treatment effect estimates measuring the treatment in 1992, 5 years before these papers switched.

Explanatory variables	All		Habitual readers	
(measured in 1992)	OLS	Probit	OLS	Probit
Treatment (1992)	0.07	0.37	0.07	0.37
	(0.03)	(0.14)	(0.03)	(0.14)
Prior Labour Vote	-0.45	1.01	-0.45	1.01
	(0.04)	(0.20)	(0.04)	(0.20)
Prior Conservative Vote	-0.22	-0.28	-0.22	-0.28
	(0.04)	(0.21)	(0.04)	(0.21)
Prior Liberal Vote	-0.03	-0.21	-0.03	-0.21
	(0.04)	(0.20)	(0.04)	(0.20)
Prior Labour Party Identification		0.94		0.94
, ,		(0.25)		(0.25)
Prior Conservative Party Identification		-0.31		-0.31
J		(0.24)		(0.24)
Prior Liberal Party Identification		0.40		0.40
,		(0.19)		(0.19)
Prior Labour Party Support		0.03		0.03
7 11		(0.19)		(0.19)
Prior Conservative Party Support		0.33		0.33
5 11		(0.19)		(0.19)
Prior Middle Political Knowledge		-0.21		-0.21
		(0.13)		(0.13)
Prior High Political Knowledge		-0.36		-0.36
6 6		(0.14)		(0.14)
Prior Television Viewer		0.13		0.13
		(0.10)		(0.10)
Prior Daily Newspaper Reader		-0.16		-0.16
		(0.10)		(0.10)
Prior Ideology		1.00		1.00
		(0.60)		(0.60)
Prior Ideological Moderation		0.62		0.62
<u> </u>		(0.46)		(0.46)
Prior Authoritarianism		-0.11		-0.11
		(0.38)		(0.38)
Prior Trade Union Member		0.01		0.01

	(0.11)	(0.11)
Prior Working-Class Identification	0.15	0.15
	(0.10)	(0.10)
Parents Voted Labour	0.20	0.20
	(0.10)	(0.10)
Prior Coping with Mortgage	-0.19	-0.19
	(0.09)	(0.09)
Prior Education: O Level	0.13	0.13
or Equivalent	(0.18)	(0.18)
Prior Education: A Level	0.05	0.05
or Equivalent	(0.20)	(0.20)
Prior Education: Some Higher	-0.02	-0.02
Education	(0.19)	(0.19)
Prior Education: College Degree	0.10	0.10
	(0.19)	(0.19)
Prior Income: £6000-£11,999	0.20	0.20
	(0.15)	(0.15)
Prior Income: £12,000-£19,999	-0.07	-0.07
	(0.16)	(0.16)
Prior Income: £20,000+	-0.05	-0.05
	(0.17)	(0.17)
Prior Income: No Income Given	-0.20	-0.20
	(0.20)	(0.20)
Prior Age: 25-34	-0.38	-0.38
	(0.21)	(0.21)
Prior Age: 35-44	-0.24	-0.24
	(0.21)	(0.21)
Prior Age: 45-54	-0.36	-0.36
	(0.22)	(0.22)
Prior Age: 55-59	-0.77	-0.77
	(0.26)	(0.26)
Prior Age: 60-64	-0.35	-0.35
	(0.26)	(0.26)
Prior Age: 65+	-0.52	-0.52
	(0.25)	(0.25)
Prior Age: not given	-0.84	-0.84
Conton	(0.45)	(0.45)
Gender	-0.16	-0.16 (0.10)
	(0.10)	. ,
White	-1.02	-1.02
	(0.33)	(0.33)
Prior Profession: Employer/	-0.04	-0.04
Business Owner	(0.27)	(0.27)
Prior Profession: Professional	-0.24	-0.24
	(0.27)	(0.27)

Prior Profession: Non-manual	-0.31	-0.31
Laborer	(0.27)	(0.27)
Prior Profession: Personal Service	-0.02	-0.02
	(0.28)	(0.28)
Prior Profession: Manual Laborer	0.21	0.21
	(0.32)	(0.32)
Prior Profession: Other	-0.47	-0.47
	(0.27)	(0.27)
Constant	-0.34	-0.34
	(0.25)	(0.25)
n	1593	1527
Region fixed effects	Х	Х

All covariates are measured in 1992 and coded to vary between 0 and 1. Where indicated, models include fixed effects for region, whose coefficients are not reported. Standard errors are in parentheses.

Table 3S: Probit Models of Vote in 1997 UK General Election for Habitual Readers

This table replicates Table 1A for habitual readers. The treatment effect estimates are unusually large in the probit models using matched samples. Given the small samples, these models appear to be underpowered, producing erratic estimates. When we drop the regional and occupation indicator variables, the estimates appear more consistent with other findings in the paper and with the linear probability models presented in Table 4S.

Explanatory variables		Preproce		
(measured in 1992 except Treatment)	All	Exact on Selected Variables	Genetic on All Variables	Treatment Instrumented with 1992 Measure
	0.80	6.17	1.86	0.13
Treatment				
(1996 or before)	(0.18)	(2.59)	(0.54)	(0.06)
Prior Labour Vote	1.01	4.89	-1.07	0.31
	(0.20)	(2.39)	(1.03)	(0.05)
Prior Conservative Vote	-0.27	3.90	-1.87	-0.09
	(0.21)	(2.43)	(0.94)	(0.05)
Prior Liberal Vote	-0.20	-3.92	-1.32	-0.07
	(0.20)	(2.65)	(1.13)	(0.05)
Prior Labour Party	0.97	1.98	2.88	0.24
	(0.25)	(1.72)	(1.17)	(0.06)
Prior Conservative Party	-0.31	3.46	0.31	-0.05
	(0.24)	(2.15)	(1.09)	(0.05)
Prior Liberal Party	0.43	-0.41	1.52	0.12
	(0.19)	(2.14)	(0.90)	(0.04)
Prior Labour Party	0.00	-9.88	-0.65	-0.01
-	(0.19)	(3.91)	(0.83)	(0.04)
Prior Conservative Party	0.33	2.42	0.99	0.08
-	(0.19)	(2.81)	(1.05)	(0.05)
Prior Middle Political	-0.20	-1.11	-0.35	-0.04
	(0.13)	(1.10)	(0.50)	(0.03)
Prior High Political	-0.34	-2.64	-2.11	-0.08
C	(0.14)	(1.68)	(0.90)	(0.03)
Prior Television Viewer	0.14	-0.69	1.82	0.02
	(0.10)	(0.91)	(0.58)	(0.02)
Prior Daily Newspaper	-0.18	-5.08	-1.86	-0.03
, <u>1</u>	(0.10)	(2.38)	(0.64)	(0.02)
Prior Ideology	1.04	-0.66	-1.33	0.12
<i>0</i> J	(0.61)	(2.85)	(2.45)	(0.11)
Prior Ideological	0.63	-4.62	-1.06	0.03

Prior Authoritarianism	-0.15	5.71	-0.29	0.01
	(0.38)	(3.73)	(1.85)	(0.08)
Prior Trade Union	0.00	-0.10	-0.49	-0.00
	(0.11)	(0.90)	(0.55)	(0.02)
Prior Working-Class	0.14	2.19	1.05	0.04
6	(0.10)	(1.40)	(0.51)	(0.02)
Parents Voted Labour	0.18	0.15	-0.31	0.04
	(0.10)	(0.86)	(0.43)	(0.02)
Prior Coping with	-0.19	0.56	-0.83	-0.01
1 0	(0.09)	(1.47)	(0.72)	(0.00)
Prior Education: O Level	0.16	-3.55	-1.43	0.03
or Equivalent	(0.18)	(2.16)	(1.05)	(0.04)
Prior Education: A Level	0.06	2.49	-1.08	0.02
or Equivalent	(0.20)	(2.28)	(1.09)	(0.04)
Prior Education: Some	-0.01	3.38	-2.50	-0.00
Education	(0.19)	(2.32)	(1.24)	(0.04)
Prior Education: College	0.11	2.95	-2.34	0.03
C	(0.20)	(2.23)	(1.31)	(0.04)
Prior Income: £6000-	0.23	0.41	0.11	0.05
	(0.15)	(1.30)	(0.57)	(0.03)
Prior Income: £12,000-	-0.04	0.61	-0.29	0.00
	(0.16)	(1.26)	(0.81)	(0.03)
Prior Income: £20,000+	-0.01	0.15	-0.75	0.01
	(0.17)	(1.03)	(0.72)	(0.04)
Prior Income: No Income	-0.17	-0.82	-0.60	-0.01
	(0.20)	(1.46)	(1.01)	(0.04)
Prior Age: 25-34	-0.42	5.20	-1.77	-0.10
	(0.22)	(3.74)	(1.51)	(0.05)
Prior Age: 35-44	-0.30	7.61	-1.48	-0.06
	(0.22)	(4.23)	(1.58)	(0.05)
Prior Age: 45-54	-0.42	7.32	-0.35	-0.10
	(0.23)	(4.30)	(1.59)	(0.05)
Prior Age: 55-59	-0.78	7.74	-1.13	-0.16
	(0.26)	(4.53)	(1.69)	(0.06)
Prior Age: 60-64	-0.39	-0.44	-1.81	-0.09
	(0.26)	(3.41)	(1.65)	(0.06)
Prior Age: 65+	-0.57	7.92	-2.19	-0.13
	(0.25)	(4.51)	(1.60)	(0.05)
Prior Age: not given	-0.80			-0.19
	(0.45)			(0.09)
Gender	-0.13	-5.43	-1.38	-0.03
	(0.10)	(2.49)	(0.71)	(0.02)
White	-1.11	-4.63		-0.24
	(0.34)	(8.33)		(0.07)
Prior Profession:	-0.23	-6.36	0.42	-0.01

Business Owner	(0.43)	(2.65)	(85.70)	(0.08)
Prior Profession:	-0.35	-6.84	-0.61	-0.03
Professional	(0.47)	(0.00)	(85.71)	(0.09)
Prior Profession: Non-	-0.21	-9.15	0.46	-0.01
Laborer	(0.42)	(2.47)	(85.70)	(0.08)
Prior Profession:	-0.18	-5.65	2.58	0.01
Personal Service	(0.46)	(3.16)	(85.71)	(0.09)
Prior Profession:	-0.20	-9.96	0.40	-0.01
Manual Laborer	(0.42)	(2.84)	(85.70)	(0.08)
Prior Profession: Other	-0.76	-9.88	-0.48	-0.12
	(0.46)	(2.99)	(85.70)	(0.09)
Constant	0.49	8.35	8.01	0.55
	(0.93)	(10.07)	(85.78)	(0.17)
Log likelihood	-	-471.036	-165.499	
n	1484	186	201	1484
Region fixed effects	Х	Х	Х	Х

All covariates are measured in 1992 and coded to vary between 0 and 1. Where indicated, models include fixed effects for region, whose coefficients not are reported. Standard errors are in parentheses.

Table 4S: Probit Models of Vote in 1997 UK General Election for Habitual Readers (excluding region and occupation dummies)

This table replicates Table 1A for habitual readers without region and occupation indicator variables. Table 2 uses the estimates from this table in its marginal effects calculations.

Explanatory variables	Preprocessed with Matching		
(measured in 1992 except Treatment)	Exact on Selected Variables	Genetic on All Variables	
Treatment (1996 or before)	1.54	1.20	
	(0.59)	(0.35)	
Prior Labour Vote	3.12	-0.58	
	(1.43)	(0.84)	
Prior Conservative Vote	1.16	-0.98	
	(1.15)	(0.72)	
Prior Liberal Vote	-0.16	-1.33	
	(1.20)	(0.84)	
Prior Labour Party Identification	0.61	2.77	
-	(1.14)	(0.94)	
Prior Conservative Party Identification	0.55	0.21	
	(1.16)	(0.89)	
Prior Liberal Party Identification	-1.44	0.64	
·	(1.28)	(0.72)	
Prior Labour Party Support	-3.34	-0.46	
	(1.24)	(0.60)	
Prior Conservative Party Support	-0.08	0.85	
7 11	(1.13)	(0.75)	
Prior Middle Political Knowledge	-0.38	-0.33	
	(0.50)	(0.39)	
Prior High Political Knowledge	-1.15	-1.59	
5 5	(0.65)	(0.60)	
Prior Television Viewer	-0.53	1.23	
	(0.55)	(0.41)	
Prior Daily Newspaper Reader	-1.02	-1.07	
	(0.65)	(0.46)	
Prior Ideology	-0.38	-1.17	
	(1.58)	(1.63)	
Prior Ideological Moderation	-1.55	-0.64	
C	(1.32)	(1.20)	
Prior Authoritarianism	1.55	-0.26	
	(1.51)	(1.28)	
Prior Trade Union Member	0.22	0.21	
	(0.44)	(0.39)	
Prior Working-Class Identification	1.00	0.69	

	(0.48)	(0.38)
Parents Voted Labour	0.24	-0.06
	(0.39)	(0.32)
Prior Coping with Mortgage	0.30	-0.62
	(0.67)	(0.52)
Prior Education: O Level	-1.52	-0.65
or Equivalent	(0.95)	(0.69)
Prior Education: A Level	0.37	-1.16
or Equivalent	(1.05)	(0.72)
Prior Education: Some Higher	-0.16	-1.64
Education	(0.85)	(0.76)
Prior Education: College Degree	0.34	-1.27
	(0.91)	(0.80)
Prior Income: £6000-£11,999	-0.44	0.23
	(0.58)	(0.43)
Prior Income: £12,000-£19,999	-0.21	-0.13
	(0.57)	(0.51)
Prior Income: £20,000+	-0.14	-0.51
	(0.60)	(0.50)
Prior Income: No Income Given	-0.45	0.01
	(0.74)	(0.67)
Prior Age: 25-34	0.50	-0.66
	(0.93)	(1.09)
Prior Age: 35-44	1.51	-0.34
	(0.98)	(1.09)
Prior Age: 45-54	0.96	-0.12
	(1.01)	(1.11)
Prior Age: 55-59	1.01	-0.18
	(1.54)	(1.19)
Prior Age: 60-64	-1.53	-0.42
	(1.34)	(1.21)
Prior Age: 65+	1.12	-1.25
	(1.17)	(1.13)
Gender	-1.94	-0.81
	(0.58)	(0.37)
White	-2.06	
	(1.51)	
Constant	2.56	3.61
	(2.78)	(2.44)
Log likelihood	-45.21	-66.09
n	186	201
Region fixed effects	Х	Х

All covariates are measured in 1992 and coded to vary between 0 and 1. Where indicated, models include fixed effects for region, whose coefficients not reported. Standard errors are in parentheses.

Table 5S: Additional Instrumental Variables Estimates with 1992 Baseline for All Treated

The first column of this table replicates instrumental variable analysis used in Table 2, instrumenting the treatment measured in 1996 (reading switching paper) with 1992 readership. The second column instead uses the treatment measured in 1997 (after the papers switched), instrumented with 1992 readership.

<i>Explanatory variables</i> (measured in 1992 except Treatment)	1996 Treatment Instrumented with 1992 Measure	1997 Treatment Instrumented with 1992 Measure
Treatment	0.11	0.11
	(0.05)	(0.05)
Prior Labour Vote	0.28	0.29
The Labour Vole	(0.05)	(0.05)
Prior Conservative Vote	-0.07	-0.08
	(0.05)	(0.05)
Prior Liberal Vote	-0.07	-0.06
The Liberar vote	(0.05)	(0.05)
Prior Labour Party Identification	0.22	0.24
Thor Eucour Furty Rechtmention	(0.05)	(0.06)
Prior Conservative Party Identification	-0.06	-0.05
Thoreonservative Furty Identification	(0.05)	(0.05)
Prior Liberal Party Identification	0.16	0.14
Thor Elberarr arty reentification	(0.04)	(0.04)
Prior Labour Party Support	-0.00	0.00
Thor Labour Farty Support	(0.04)	(0.04)
Prior Conservative Party Support	0.08	0.08
The conservative raity support	(0.04)	(0.04)
Prior Middle Political Knowledge	-0.06	-0.05
The made fonded knowledge	(0.03)	(0.03)
Prior High Political Knowledge	-0.08	-0.08
The fight onlice the wiedge	(0.03)	(0.03)
Prior Television Viewer	0.02	0.01
	(0.02)	(0.02)
Prior Daily Newspaper Reader	-0.03	-0.03
Thor Dury Teenspaper Reader	(0.02)	(0.02)
Prior Ideology	0.14	0.17
1110114001055	(0.11)	(0.11)
Prior Ideological Moderation	0.01	0.04
	(0.08)	(0.08)
Prior Authoritarianism	-0.00	0.02
	(0.08)	(0.08)
Prior Trade Union Member	0.01	-0.00

	(0,02)	(0,02)
Driver Working Class Identification	(0.02) 0.04	(0.02) 0.03
Prior Working-Class Identification	(0.02)	(0.03)
Parents Voted Labour	0.03	0.05
Falents voled Labour	(0.02)	(0.02)
Prior Coping with Mortgage	-0.03	-0.03
Prior Coping with Mortgage	(0.03)	(0.03)
Prior Education: O Level	0.02	0.03
or Equivalent	(0.04)	(0.04)
Prior Education: A Level	0.02	0.02
or Equivalent	(0.04)	(0.02)
Prior Education: Some Higher	-0.01	-0.01
Education	(0.04)	(0.04)
Prior Education: College Degree	0.02	0.03
Filor Education. Conege Degree	(0.04)	(0.04)
Prior Income: £6000-£11,999	0.07	0.06
11101 Income. 20000-211,999	(0.03)	(0.03)
Prior Income: £12,000-£19,999	0.00	0.01
11101 meome. £12,000-£19,999	(0.03)	(0.03)
Prior Income: £20,000+	0.00	0.00
1 nor meone. 220,000+	(0.04)	(0.04)
Prior Income: No Income Given	-0.00	-0.00
Thor meenie. No meenie Green	(0.04)	(0.04)
Prior Age: 25-34	-0.10	-0.11
111011150.25 51	(0.04)	(0.05)
Prior Age: 35-44	-0.07	-0.08
1.101.1160.00	(0.04)	(0.05)
Prior Age: 45-54	-0.11	-0.11
	(0.05)	(0.05)
Prior Age: 55-59	-0.16	-0.16
6	(0.05)	(0.05)
Prior Age: 60-64	-0.09	-0.10
e	(0.05)	(0.06)
Prior Age: 65+	-0.12	-0.13
C	(0.05)	(0.05)
Prior Age: not given	-0.17	-0.20
0 0	(0.09)	(0.09)
Gender	-0.03	-0.03
	(0.02)	(0.02)
White	-0.20	-0.23
	(0.06)	(0.07)
Prior Profession: Employer/	0.00	0.00
Business Owner	(0.08)	(0.08)
Prior Profession: Professional	-0.03	-0.02
	(0.09)	(0.09)
	(0.09)	(0.07)

Prior Profession: Non-manual	0.01	0.01
Laborer	(0.07)	(0.08)
Prior Profession: Personal Service	0.03	0.02
	(0.08)	(0.09)
Prior Profession: Manual Laborer	0.00	-0.00
	(0.08)	(0.08)
Prior Profession: Other	-0.11	-0.10
	(0.08)	(0.09)
Constant	0.53	0.51
	(0.17)	(0.17)
n	1593	1527
Region fixed effects	Х	Х

All covariates are measured in 1992 and coded to vary between 0 and 1. Where indicated, models include fixed effects for region, whose coefficients not reported. Standard errors are in parentheses.

Table 6S: Replication of Table 1A with 1996 Baseline

This table replicates Table 1A, except with all explanatory variables measured (if possible) in 1996 rather than 1992. Those variables not asked in 1996 are measured in 1992. Using a 1996 baseline allows us to include two additional explanatory variables that are plausibly important not available in 1992 wave: perceptions of the *1996 Economy* and *Prior European Integration Views*. The model in the second column includes these as additional control variables.

<i>Explanatory variables</i> (Measured in 1996 when possible)	Same Variables as in Table 1A	With 1996 Economy and European Integration Views
Treatment	0.71	0.73
	(0.17)	(0.17)
Prior Labour Vote Intention	0.87	0.90
(1996 wave)	(0.19)	(0.19)
Prior Conservative Vote Intention	-0.82	-0.81
(1996 wave)	(0.27)	(0.27)
Prior Liberal Vote Intention	-0.40	-0.41
(1996 wave)	(0.21)	(0.22)
Prior Labour Party Identification	1.40	1.42
(1996 wave)	(0.35)	(0.36)
Prior Conservative Party Identification	0.05	0.20
(1996 wave)	(0.31)	(0.32)
Prior Liberal Party Identification	0.55	0.55
(1996 wave)	(0.20)	(0.20)
Prior Labour Party Support	-0.31	-0.23
(1996 wave)	(0.24)	(0.24)
Prior Conservative Party Support	-0.20	-0.13
(1996 wave)	(0.23)	(0.24)
Prior Middle Political Knowledge	-0.03	-0.04
(1992 wave)	(0.16)	(0.16)
Prior High Political Knowledge	-0.12	-0.13
(1992 wave)	(0.18)	(0.18)
Prior Television Viewer	0.10	0.08
(1992 wave)	(0.12)	(0.12)
Prior Daily Newspaper Reader	-0.41	-0.43
(1996 wave)	(0.12)	(0.13)
Prior Ideology	0.74	0.62
(1996 wave)	(0.59)	(0.60)
Prior Ideological Moderation	0.14	0.16
(1996 wave)	(0.37)	(0.38)
Prior Authoritarianism	0.47	0.31
(1996 wave)	(0.42)	(0.44)
Prior Trade Union Member	0.09	0.09
(1992 wave)	(0.13)	(0.13)
Prior Working-Class Identification	0.20	0.21

(1992 wave)	(0.12)	(0.13)
Parents Voted Labour	0.30	0.32
(1992 wave)	(0.11)	(0.12)
Prior Coping with Mortgage	-0.09	-0.10
(1995 wave)	(0.08)	(0.08)
Prior Education: O Level	0.00	0.05
or Equivalent	(0.22)	(0.22)
Prior Education: A Level	-0.04	0.00
or Equivalent	(0.24)	(0.24)
Prior Education: Some Higher	-0.14	-0.10
Education	(0.23) 0.07	(0.23)
Prior Education: College Degree (1992 wave)		0.15
	(0.23) 0.26	(0.24) 0.28
Prior Income: £6000-£11,999		(0.19)
(1992 wave)	(0.19) -0.12	-0.13
Prior Income: £12,000-£19,999	(0.20)	(0.20)
(1992 wave)	-0.18	-0.18
Prior Income: £20,000+	(0.20)	(0.20)
(1992 wave) Prior Income: No Income Given	-0.35	-0.36
(1992 wave)	(0.25)	(0.25)
Prior Age: 25-34	0.04	0.06
(1992 wave)	(0.24)	(0.25)
Prior Age: 35-44	0.17	0.20
(1992 wave)	(0.24)	(0.25)
Prior Age: 45-54	0.10	0.09
(1992 wave)	(0.26)	(0.26)
Prior Age: 55-59	0.01	-0.07
(1992 wave)	(0.30)	(0.31)
Prior Age: 60-64	0.07	0.10
(1992 wave)	(0.31)	(0.31)
Prior Age: 65+	0.21	0.20
(1992 wave)	(0.29)	(0.29)
Prior Age: not given	-0.13	-0.13
(1992 wave)	(0.50)	(0.50)
Gender	0.04	0.01
(1992 wave)	(0.12)	(0.13)
White	-0.47	-0.48
(1992 wave)	(0.38)	(0.39)
Prior Profession: Employer/	-0.34	-0.31
Business Owner (1992 wave)	(0.46)	(0.47)
Prior Profession: Professional	-0.56	-0.51
(1992 wave)	(0.50)	(0.51)
Prior Profession: Non-manual	-0.37	-0.33
i noi i nicoston. non-manual	-0.57	-0.55

Laborer (1992 wave)	(0.44)	(0.44)
Prior Profession: Personal Service	-0.49	-0.48
(1992 wave)	(0.49)	(0.50)
Prior Profession: Manual Laborer	-0.35	-0.34
(1992 wave)	(0.45)	(0.46)
Prior Profession: Other	-0.94	-0.98
(1992 wave)	(0.50)	(0.51)
1996 Economy		0.74
(1996 wave)		(0.27)
Prior European Integration Views		0.46
(1996 wave)		(0.17)
Constant	-1.21	-1.68
	(0.98)	(1.01)
Log likelihood	-387.8	-380.5
n	1330	1330
Region fixed effects	Х	Х

All covariates vary between 0 and 1. Where indicated, models include fixed effects for region, whose coefficients not reported. Standard errors are in parentheses.

Disaggregating the Control Group

Figure 2S: Comparing the Treated with Four Control Groups



This figure shows that the treatment effect persists when comparing treated respondents to various control groups, including (persistent) Conservative paper readers, Labour papers readers, other or no affiliation paper readers, and those who did not read newspapers. To facilitate comparisons, we mean-difference the control groups in 1996, ensuring the treatment group and each control group have the same mean in 1996. Readership is measured in 1996. Respondents who failed to report a vote choice or vote intent in any of the three waves are excluded from the analysis, resulting in smaller *n*s. Confidence intervals show plus or minus one standard error.

Table 7S: The Effect of Reading Switching Papers and Papers with Consistent Partisan Loyalties

This table replicates Table 1A. However, unlike Table 1A, where the treatment is a simple variable indicating whether the respondent read a switching paper or not, here we measure readership with more detail. The model here includes indicator variables for reading a paper that *Switched to Labour*, reading a *Faithful Conservative* paper, reading a *Faithful Labour* paper, or reading a paper that does not fit in these categories, which is labeled *Other*. The omitted category is not reading a newspaper at all.)

Explanatory variables (measured in 1992 except Treatment)	All (DV: Vote Labour in 1997)
Newspaper read in 1996	
Switched to Labour	0.57
	(0.14)
Faithful Conservative	-0.06
	(0.13)
Faithful Labour	0.48
	(0.14)
Other	-0.39
	(0.18)
Prior Labour Vote	0.86
	(0.19)
Prior Conservative Vote	-0.19
	(0.20)
Prior Liberal Vote	-0.22
	(0.19)
Prior Labour Party Identification	0.91
	(0.24)
Prior Conservative Party Identification	-0.31
	(0.23)
Prior Liberal Party Identification	0.54
	(0.18)
Prior Labour Party Support	0.04
	(0.18)
Prior Conservative Party Support	0.34
	(0.18)
Prior Middle Political Knowledge	-0.21
	(0.12)
Prior High Political Knowledge	-0.30
c c	(0.14)
Prior Television Viewer	0.16
	(0.10)
Prior Daily Newspaper Reader	-0.26
× 1 1	(0.11)
Prior Ideology	1.05

	$(0, \overline{0})$
Drive Idealagical Madagation	(0.59) 0.54
Prior Ideological Moderation	(0.44)
Prior Authoritarianism	-0.13
Filor Authoritarianism	(0.36)
Prior Trade Union Member	0.02
Thoi Hade onion Memoer	(0.11)
Prior Working-Class Identification	0.18
	(0.10)
Parents Voted Labour	0.12
	(0.09)
Prior Coping with Mortgage	-0.08
	(0.13)
Prior Education: O Level	0.12
or Equivalent	(0.18)
Prior Education: A Level	0.09
or Equivalent	(0.20)
Prior Education: Some Higher	-0.07
Education	(0.19)
Prior Education: College Degree	0.08
	(0.19)
Prior Income: £6000-£11,999	0.32
	(0.15)
Prior Income: £12,000-£19,999	-0.04
D: 1 (20.000	(0.16)
Prior Income: £20,000+	0.03
Prior Income: No Income Given	(0.16) -0.08
Filor income. No income Given	-0.08 (0.19)
Prior Age: 25-34	-0.40
Thoi Age. 23-34	(0.20)
Prior Age: 35-44	-0.24
	(0.20)
Prior Age: 45-54	-0.41
	(0.21)
Prior Age: 55-59	-0.76
	(0.25)
Prior Age: 60-64	-0.34
	(0.25)
Prior Age: 65+	-0.57
	(0.23)
Prior Age: not given	-0.72
	(0.45)
Gender	-0.12
	(0.10)

White	-0.91
	(0.30)
Prior Profession: Employer/	-0.11
Business Owner	(0.37)
Prior Profession: Professional	-0.29
	(0.42)
Prior Profession: Non-manual	-0.09
Laborer	(0.36)
Prior Profession: Personal Service	-0.05
	(0.40)
Prior Profession: Manual Laborer	-0.13
	(0.37)
Prior Profession: Other	-0.59
	(0.41)
Constant	0.00
	(0.88)
Log likelihood	-594.0
n	1593
Region fixed effects	Х

All covariates vary between 0 and 1. Includes fixed effects for region, with coefficients not reported. Standard errors are in parentheses.

Panel Attrition

As we discuss in the paper, one might worry that strong conservative supporters would drop out of the panel because they would rather not talk about politics in 1997. If so, the remaining individuals in the panel would be potentially more susceptible to persuasion. In fact, however, strong Conservative supporters are less likely, not more likely, to drop out. 1992 Labour voters are about 3% more likely to drop out than 1992 Conservative voters. On Prior Ideology, moderates are most likely to drop out, followed by liberals, then by conservatives.

The figure below shows that individuals with the highest Prior Labour Party Support (measured in 1992, see Table 1 in the paper for details) dropped out of the panel at a higher rate than did those with the lowest support: 50% versus 43%.





Further Addressing Concerns about Self-Selection

In the paper, we address concerns about self-selection between the 1992 and 1996 waves. Concerns about self-selection are also mitigated by the fact that panel members do not appear to have foreseen the switch in advance. In every panel wave, respondents were asked which party their newspaper favored. In general, Conservative paper readers saw their papers shift against the Conservative Party between 1992 and 1996. For the *Sun*, this is a sizable shift: 80% and 51% of its readers thought they favored the Conservative papers. After this initial drop, however, perceptions remained stable. In the *Sun*'s case, 47% thought they supported the Conservatives in 1995 and 57% in 1996. In a manipulation check, we find that in their post 1997 election interviews, after the *Sun* had switched, only 2% still thought it backed the Conservatives.

Table 8S: Change in Readership by Paper Read in 1992

This table shows the patterns of newspaper readership discussed in the paper. We see no evidence that 1992 Conservative voters shifted away from switching papers between 1992 and 1996 (when we measure readership). Only after 1996 do we do see some evidence of this behavior. Thus, self-selection seems unlikely to bias estimates because we measure the treatment in 1996.

Readers of switching paper in 1992			
All	Readership in 1992	Readership in 1996	Readership in 1997
Read Tory Paper	0.0	5.7	7.8
Read Labour Paper	0.0	7.4	4.8
Read Switching Paper	100.0	63.5	57.4
Read the Times	0.0	2.6	3.5
Did Not Read	0.0	18.3	15.2
Read Other	0.0	2.6	11.3
n	230	230	230
Voted for Labour in 1992	Readership in 1992	Readership in 1996	Readership in 1997
Read Tory Paper	0.0	1.4	1.4
Pand Labour Dapar	0.0	11.3	12

0.0	1.4	1.4
0.0	11.3	4.2
100.0	63.4	63.4
0.0	2.8	2.8
0.0	18.3	16.9
0.0	2.8	11.3
71	71	71
	0.0 100.0 0.0 0.0	$\begin{array}{ccc} 0.0 & 11.3 \\ 100.0 & 63.4 \\ 0.0 & 2.8 \\ 0.0 & 18.3 \end{array}$

Voted for the Tories in 1992	Readership in 1992	Readership in 1996	Readership in 1997
Read Tory Paper	0.0	9.4	11.8
Read Labour Paper	0.0	3.5	3.5
Read Switching Paper	100.0	63.5	50.6
Read the Times	0.0	4.7	5.9
Did Not Read	0.0	17.7	17.7
Read Other	0.0	1.2	10.6
n	85	85	85

Readers of Tory papers in 1992

	Readership in 1992	Readership in 1996	Readership in 1997
Read Tory Paper	100.0	68.2	67.0
Read Labour Paper	0.0	4.7	3.1
Read Switching Paper	0.0	5.0	3.4
Read the Times	0.0	3.1	3.1
Did Not Read	0.0	15.9	14.6
Read Other	0.0	3.1	8.7
n	321	321	321

Readers of Labour papers in 1992			
	Readership in 1992	Readership in 1996	Readership in 1997
Read Tory Paper	0.0	7.1	7.1
Read Labour Paper	100.0	70.5	62.7
Read Switching Paper	0.0	7.6	6.6
Read the Times	0.0	1.0	0.5
Did Not Read	0.0	11.0	13.2
Read Other	0.0	2.9	10.0
n	410	410	410

"What Can Go Wrong" with Matching

Reviewer #2 asked for information about what can go wrong with matching. Here, we excerpt a discussion on this issue from the following paper and then discuss briefly why we avoid the potential problems they raise.

Ho, Daniel E., Kosuke Imai, Gary King, and Elizabeth A. Stuart. 2007. "Matching as Nonparametric Preprocessing for Reducing Model Dependence in Parametric Causal Inference." *Political Analysis* 15 (3): 199-236.

8 What Can Go Wrong

The advantage of matching is that it is relatively robust to small changes in procedures and produces a data set that is by design less sensitive to modeling assumptions. However, like any method, using it badly or to ill effect is certainly possible. Thus, in this section, we discuss four ways in which preprocessing can go wrong and how researchers might try to avoid these problems.

First, since the curse of dimensionality affects balancing diagnostics, we may well miss a higher dimensional aspect of imbalance when checking lower dimensional summaries. Even if we are uninterested in testing these with our parametric model, they can affect our estimates. Such will be the case with parametric models with or without preprocessing, and so in all but the most unusual cases preprocessing should at least not make things worse. One pathological case where preprocessing could hurt is if some covariate has a huge effect on the outcome variable and preprocessing slightly reduces balance on this variable but improves it for all the others. A researcher might be fooled into choosing a matching trade-off like this if he or she were not aware of the large effect of this covariate. Carefully evaluating what covariates are likely to have the largest effects, and using multiple measures of balance, are essential to avoid this pitfall.

Second, as with all statistical methods, a bias-variance trade-off exists for matching. If we drop many observations during preprocessing, and balance is not substantially improved, the mean squared error (or other mean-variance summary) of the estimated causal effect might actually increase. Users must pay close attention to this trade-off during the process of matching, but unfortunately no precise rules exist for how to make these choices. In particular, the methodological literature offers no formal estimates of mean squared error and so in marginal cases it can be difficult to know whether or how much preprocessing will help. Of course, dropping observations does not necessarily mean that preprocessing is worse since improving balance can also increase efficiency, and in any event including imbalanced observations requiring extrapolation in a parametric analysis

merely produces false precision. So although estimated standard errors may increase in some cases with preprocessing, they would likely be more accurate. Moreover, in many situations, eliminating observations far from the rest of the data as matching does will reduce heterogeneity and thereby further reduce variance.

Third, the matching literature offers a large number of possible and seemingly ad hoc procedures. From one perspective, we might be concerned about the sensitivity of our

results to changes in this process, just as we have been concerned with the sensitivity of causal effect estimates to parametric modeling assumptions. This is not normally viewed as a major issue since the right procedure is the one that maximizes balance (with n as large as possible), no matter how many procedures we try. By applying this criterion in a disciplined way (i.e., without consulting Y) to a large number of possible matching procedures, no choices are open to the analyst. Instead, researchers should merely run as many as possible and choose by maximizing balance. Unlike parametric modeling exercises, we need not choose this matching procedure or another; we merely run as many as feasible, particularly those most likely to reduce bias and model dependence, and apply this criterion.

Finally, by dropping observations, we may wind up losing some critically important cases or may change either the information base of our sample or, in special cases such as when dropping treated units, the definition of the causal effect. Examining the dropped cases provides an easy diagnostic for this problem. However, we must be alert to the problem that if we learn that some critical units are dropped, then it may mean that no appropriate matches can be found for them. In this situation, we may be forced to conclude that the data do not contain sufficient information to answer the questions posed, no matter what method is chosen.

Discussion of These Problems for Our Findings

We do not think that any of these problems apply to our case. In part, this is because we use matching as a robustness check. We find essentially the same results in bivariate analyses, in models that include many controls, with instrumental variables estimates, and with two approaches to matching (exact and genetic). If we only found our results in the matching estimates, we would be more concerned about potential problems with matching. We also avoid the problems they raise by taking additional steps that we now describe.

The first concern Ho et al. (2007) raise is that matching might prioritize balance on less important variables at the cost of imbalance on more important variables. We address this by specifying, in the analysis with exact and genetic matching, that prior vote, prior partisan identification, prior party approval, and political knowledge be exactly matched. As is evident in the balance table (Table 3 in the paper, see also the additional balance statistics below), the matching procedures eliminate differences between the treated and untreated groups on these variables. In this way, we ensure that the matching does not create imbalances on these most-important variables.

Another concern they raise is that imbalances may continue to exist on higher order statistics, such as the treatment and control groups having different variances on matched covariates. To address these concerns, we present numerous additional balance statistics below. These statistics suggest that the treatment and control groups are very similar after the genetic matching.

Finally, Ho et al. (2007) also discuss potential problems with dropping observations. In our case, we have so many more potential control cases that dropping unmatched control cases only brings benefits. By dropping them, we are eliminating individuals in the control group who are very different from treated individuals on important characteristics. Because they are so different, comparing the treated with these control individuals could lead to incorrect inferences.

Additional Balance Statistics for Genetic Matching

For details on these balance statistics see:

- Sekhon, Jasjeet S. 2007a. "Alternative Balance Metrics for Bias Reduction in Matching Methods for Causal Inference." University of California, Berkeley. Typescript.
- Sekhon, Jasjeet S. 2007b. "Matching." Version 4.3-1. University of California, Berkeley, <u>http://sekhon.berkeley.edu</u>.
- Sekhon, Jasjeet S. Forthcoming. "Multivariate and Propensity Score Matching Software with Automated Balance Optimization." *Journal of Statistical Software*.

Prior Labour Vote

	Before Matching	After Matching
mean treatment	0.38863	0.38863
mean control	0.32272	0.38863
std mean diff	13.489	0
mean raw eQQ diff	0.066351	0
med raw eQQ diff	0	0
max raw eQQ diff	1	0
mean eCDF diff	0.032952	0
med eCDF diff	0.032952	0
max eCDF diff	0.065905	0
var ratio (Tr/Co)	1.0914	1
T-test p-value	0.067572	1

Prior Conservative Vote

	Before Matching	After Matching
mean treatment	0.38863	0.38863
mean control	0.40376	0.38863
std mean diff	-3.0981	0
mean raw eQQ diff	0.014218	0
med raw eQQ diff	0	0
max raw eQQ diff	1	0
mean eCDF diff	0.0075685	0
med eCDF diff	0.0075685	0
max eCDF diff	0.015137	0
	0.99093	1
T-test p-value	0.67561	1

Prior Liberal Vote

	Before Matching	After Matching
mean treatment	0.15640	0.15640
mean control	0.18813	0.15640
std mean diff	-8.7161	0
mean raw eQQ diff	0.033175	0
med raw eQQ diff	0	0
max raw eQQ diff	1	0
mean eCDF diff	0.015868	0
med eCDF diff	0.015868	0
max eCDF diff	0.031735	0
var ratio (Tr/Co)	0.8673	1
T-test p-value	0.24397	1

Prior Labour Party Support

	Before Matching	After Matching
mean treatment	0.48717	0.48717
mean control	0.46235	0.49171
std mean diff	7.578	-1.3855
mean raw eQQ diff	0.024653	0.014419
med raw eQQ diff	0	0
max raw eQQ diff	0.25	0.25
mean eCDF diff	0.018503	0.012638
med eCDF diff	0.012172	0.011848
max eCDF diff	0.047561	0.028436
var ratio (Tr/Co)	1.0230	1.0768
T-test p-value	0.30541	0.68282
KS Bootstrap p-value	0.4012	0.9144
KS Naive p-value	0.80202	1
KS Statistic	0.047561	0.028436

Prior Labour Party Identification

	Before Matching	After Matching
mean treatment	0.33649	0.33649
mean control	0.31476	0.34123
std mean diff	4.5883	-1.0006
mean raw eQQ diff	0.023697	0.0047393
med raw eQQ diff	0	0
max raw eQQ diff	1	1
mean eCDF diff	0.010866	0.0023697
med eCDF diff	0.010866	0.0023697
max eCDF diff	0.021732	0.0047393
var ratio (Tr/Co)	1.0393	0.9932
T-test p-value	0.53423	0.70574

Prior Conservative Party Identification

	Before Matching	After Matching
mean treatment	0.41232	0.41232
mean control	0.41896	0.40758
std mean diff	-1.3448	0.9605
mean raw eQQ diff	0.0047393	0.0047393
med raw eQQ diff	0	0
max raw eQQ diff	1	1
mean eCDF diff	0.0033179	0.0023697
med eCDF diff	0.0033179	0.0023697
max eCDF diff	0.0066358	0.0047393
var ratio (Tr/Co)	0.99942	1.0035
T-test p-value	0.85576	0.78175

Prior Liberal Party Identification

<i>.</i>	Before Matching	After Matching
mean treatment	0.13270	0.13270
mean control	0.15412	0.12796
std mean diff	-6.2998	1.3937
mean raw eQQ diff	0.023697	0.0047393
med raw eQQ diff	0	0
max raw eQQ diff	1	1
mean eCDF diff	0.010712	0.0023697
med eCDF diff	0.010712	0.0023697
max eCDF diff	0.021423	0.0047393
var ratio (Tr/Co)	0.88637	1.0314
T-test p-value	0.3987	0.70574

Prior Conservative Party Support

· · · ·	Before Matching	After Matching
mean treatment	0.5237	0.5237
mean control	0.52058	0.51896
std mean diff	0.93566	1.4241
mean raw eQQ diff	0.0061025	0.023697
med raw eQQ diff	0	0
max raw eQQ diff	0.25	0.25
mean eCDF diff	0.0075828	0.018957
med eCDF diff	0.0078703	0.014218
max eCDF diff	0.012603	0.042654
var ratio (Tr/Co)	0.9981	1.1181
T-test p-value	0.89937	0.72395
KS Bootstrap p-value	0.99	0.7068
KS Naive p-value	1	0.99075
KS Statistic	0.012603	0.042654

Prior Political Knowledge

	Before Matching	After Matching
mean treatment	0.54502	0.54502
mean control	0.67113	0.55213
std mean diff	-32.916	-1.8556
mean raw eQQ diff	0.12559	0.021327
med raw eQQ diff	0	0
max raw eQQ diff	0.5	0.5
mean eCDF diff	0.08407	0.014218
med eCDF diff	0.10719	0.014218
max eCDF diff	0.14502	0.028436
var ratio (Tr/Co)	1.1434	1.0843
T-test p-value	1.0532e-05	0.60182
KS Bootstrap p-value	2e-04	0.7874
KS Naive p-value	0.00090618	1
KS Statistic	0.14502	0.028436

Prior Working-Class Identification

0 0	Before Matching	After Matching
mean treatment	0.71564	0.71564
mean control	0.58123	0.71677
std mean diff	29.724	-0.24933
mean raw eQQ diff	0.13473	0.0043633
med raw eQQ diff		0
max raw eQQ diff	1	0.57927
mean eCDF diff	0.11216	0.0023697
med eCDF diff	0.13424	0.0023697
max eCDF diff	0.13605	0.0047393
var ratio (Tr/Co)	0.84184	1.0135
T-test p-value	8.9685e-05	0.88309
KS Bootstrap p-value	< 2.22e-16	0.9894
KS Naive p-value		1
KS Statistic	0.13605	0.0047393

Parents Voted Labour

	Before Matching	After Matching
mean treatment	0.43602	0.43602
mean control	0.35384	0.43602
std mean diff	16.534	0
mean raw eQQ diff	0.080569	0
med raw eQQ diff	0	0
max raw eQQ diff	1	0
mean eCDF diff	0.041092	0
med eCDF diff	0.041092	0
max eCDF diff	0.082184	0
var ratio (Tr/Co)	1.0799	1
T-test p-value	0.025375	1

Prior Television Viewer

	Before Matching	After Matching
mean treatment	0.21801	0.21801
mean control	0.28944	0.20853
std mean diff	-17.258	2.2902
mean raw eQQ diff	0.07109	0.0094787
med raw eQQ diff	0	0
max raw eQQ diff	1	1
mean eCDF diff	0.035713	0.0047393
med eCDF diff	0.035713	0.0047393
max eCDF diff	0.071426	0.0094787
var ratio (Tr/Co)	0.83228	1.0329
T-test p-value	0.021901	0.56401

Prior Ideology

	Before Matching	After Matching
mean treatment	0.54993	0.54993
mean control	0.535	0.54842
std mean diff	8.1908	0.8274
mean raw eQQ diff	0.021349	0.020957
med raw eQQ diff	0	0
max raw eQQ diff	0.09091	0.09091
mean eCDF diff	0.026593	0.022860
med eCDF diff	0.027951	0.018957
max eCDF diff	0.052945	0.056872
var ratio (Tr/Co)	0.85258	1.3228
T-test p-value	0.27411	0.85725
KS Bootstrap p-value	0.4626	0.6974
KS Naive p-value	0.6839	0.88455
KS Statistic	0.052945	0.056872

Prior Authoritarianism

	Before Matching	After Matching
mean treatment	0.5789	0.5789
mean control	0.57032	0.57254
std mean diff	5.6916	4.2149
mean raw eQQ diff	0.014950	0.027072
med raw eQQ diff	0	0.04
max raw eQQ diff	0.12	0.24
mean eCDF diff	0.023224	0.032731
med eCDF diff	0.019307	0.035545
max eCDF diff	0.049633	0.07583
var ratio (Tr/Co)	1.2888	1.5628
T-test p-value	0.43507	0.50069
KS Bootstrap p-value	0.5124	0.3648
KS Naive p-value	0.75798	0.57888
KS Statistic	0.049633	0.07583

Prior Trade Union Member

	Before Matching	After Matching
mean treatment	0.21801	0.21801
mean control	0.24023	0.19905
std mean diff	-5.3693	4.5804
mean raw eQQ diff	0.023697	0.018957
med raw eQQ diff	0	0
max raw eQQ diff	1	1
mean eCDF diff	0.011111	0.0094787
med eCDF diff	0.011111	0.0094787
max eCDF diff	0.022222	0.018957
var ratio (Tr/Co)	0.9378	1.0693
T-test p-value	0.47011	0.61739

Prior Coping with Mortgage

The copies will mongage		
	Before Matching	After Matching
mean treatment	0.70853	0.70853
mean control	0.66218	0.6909
std mean diff	12.62	4.799
mean raw eQQ diff	0.047393	0.017625
med raw eQQ diff	0	0
<pre>max raw eQQ diff</pre>	0.5	0.5
mean eCDF diff	0.047955	0.013033
med eCDF diff	0.064343	0.016588
max eCDF diff	0.066152	0.018957
var ratio (Tr/Co)	0.95013	0.94766
T-test p-value	0.089945	0.57228
KS Bootstrap p-value	0.0924	0.8952
KS Naive p-value	0.39966	1
KS Statistic	0.066152	0.018957

Prior Education

	Before Matching	After Matching
mean treatment	0.74763	0.74763
mean control	0.64237	0.73934
std mean diff	32.804	2.5846
mean raw eQQ diff	0.10664	0.0082938
med raw eQQ diff	0	0
max raw eQQ diff	0.5	0.25
mean eCDF diff	0.084211	0.0066351
med eCDF diff	0.11887	0
max eCDF diff	0.14158	0.028436
var ratio (Tr/Co)	0.80773	1.0158
T-test p-value	1.7235e-05	0.59702
KS Bootstrap p-value	< 2.22e-16	0.8844
KS Naive p-value	0.0013000	1
KS Statistic	0.14158	0.028436

Prior Income

	Before Matching	After Matching
mean treatment	0.48717	0.48717
mean control	0.56077	0.49605
std mean diff	-21.562	-2.6013
mean raw eQQ diff	0.073308	0.015718
med raw eQQ diff	0	0
max raw eQQ diff	0.30454	0.30454
mean eCDF diff	0.075688	0.011058
med eCDF diff	0.076589	0.0094787
max eCDF diff	0.092249	0.042654
var ratio (Tr/Co)	1.1196	1.0773
T-test p-value	0.0035862	0.53027
KS Bootstrap p-value	0.032	0.8098
KS Naive p-value	0.088705	0.99075
KS Statistic	0.092249	0.042654

Prior Age

Before Matching	After Matching
0.46922	0.46922
0.48763	0.46634
-5.9425	0.93056
0.022346	0.019119
0	0
0.16667	0.16667
0.034690	0.011301
0.043534	0.0047393
0.051855	0.047393
1.0648	1.1681
0.42022	0.7886
0.3556	0.725
0.7086	0.97177
0.051855	0.047393
	0.46922 0.48763 -5.9425 0.022346 0 0.16667 0.034690 0.043534 0.051855 1.0648 0.42022 0.3556 0.7086

Gender

	Before Matching	After Matching
mean treatment	0.45498	0.45498
mean control	0.55572	0.45498
std mean diff	-20.182	0
mean raw eQQ diff	0.099526	0
med raw eQQ diff	0	0
<pre>max raw eQQ diff</pre>	1	0
mean eCDF diff	0.05037	0
med eCDF diff	0.05037	0
max eCDF diff	0.10074	0
var ratio (Tr/Co)		1
T-test p-value	0.0066971	1

White

	Before Matching	After Matching
mean treatment	0.98578	0.98578
mean control	0.97612	0.98578
std mean diff	8.1406	0
mean raw eQQ diff	0.014218	0
med raw eQQ diff	0	0
max raw eQQ diff	1	0
mean eCDF diff	0.0048302	0
med eCDF diff	0.0048302	0
max eCDF diff	0.0096604	0
var ratio (Tr/Co)	0.60375	1
T-test p-value	0.29155	1

Before Matching Minimum p.value: < 2.22e-16 Variable Names: Xwkclass Xhedqul92

After Matching Minimum p.value: 0.3648 Variable Name: Xauth92